

Note on Computing Ratings from Eigenvectors *

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In memory of
Gene Howard Golub
1932–2007

Abstract

We consider the problem of computing ratings using the results of games played between a set of n players, and show how this problem can be reduced to computing the positive eigenvectors corresponding to the dominant eigenvalues of certain $n \times n$ matrices. There is a close connection with the stationary probability distributions of certain Markov chains. In practice, if n is large, then the matrices involved will be sparse, and the power method may be used to solve the eigenvalue problems efficiently. We give an algorithm based on the power method, and also derive the same algorithm by an independent method.

1 Introduction

Suppose that n players, numbered $1, \dots, n$, play a total of m games which may end in a win, loss or draw. We assume that a win scores 1 point, a draw 0.5, and a loss 0. The results can be summarised by an $n \times n$ *score matrix* $S = (s_{i,j})$, where $s_{i,j}$ is the number of points that player i scores against player j ($i \neq j$). The diagonal entries $s_{i,i}$ are arbitrary; for reasons discussed later we assume that $s_{i,i} = \sigma$, where $\sigma \geq 0$ is a constant. For the moment, the reader may assume that $\sigma = 0$.

The aim is to assign *ratings* r_i to the players in such a way that players who perform better obtain higher ratings. The problem may arise when S represents the results obtained in a single event, e.g. a Swiss tournament, and typically we want to use the ratings to break ties between players on equal scores. It may also arise when S represents all the recent results involving a large set of players, for example in the regular updates to a national rating system. In the latter case, we can expect the matrix S to be large and sparse.

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The expected score of one player when playing against another should depend only on the difference of their ratings. Thus, we assume that, if the ratings are correct, then the expected score of player i in a game against player j is $f(r_i - r_j)$, for some function $f : \mathcal{R} \rightarrow [0, 1]$. Since the total score obtained by both players in a game is 1, the function f should satisfy the condition

$$f(z) + f(-z) = 1 . \tag{1}$$

It is reasonable to assume that $f(z)$ is monotonic increasing, and that

$$0 = \lim_{y \rightarrow -\infty} f(y) < f(z) < \lim_{y \rightarrow +\infty} f(y) = 1 . \tag{2}$$

It is easy to find functions satisfying these conditions. For example, if $\phi(x)$ is a probability density on $(-\infty, +\infty)$, satisfying the condition $\phi(x) > 0$ and $\phi(x) = \phi(-x)$, then

$$f(z) = \int_{-\infty}^z \phi(x) dx$$

satisfies (1) and (2). We could take the normal probability density

$$\phi(x) = \frac{1}{\sqrt{2\pi}} \exp(-x^2/2) ,$$

but there seems to be no real justification for this choice, and we shall make a more convenient choice below.

Let $g(z) = f(z)/(1 - f(z))$. For a game without draws, $g(r_i - r_j)$ is the ratio $Prob(\text{player } i \text{ wins})/Prob(\text{player } j \text{ wins})$. From (1), $g(z) = f(z)/f(-z)$ and

$$g(z)g(-z) = 1 . \tag{3}$$

A simple solution to (3) is $g(z) = e^{cz}$ for some constant c . Since $f(z) = g(z)/(1 + g(z))$, this suggests taking

$$f(z) = \frac{1}{1 + e^{-cz}} . \tag{4}$$

It is easy to check that f satisfies (1) and (2) if $c > 0$.

We shall *assume* that $f(z)$ has the form (4). This involves an empirical assumption that could be tested by experiment. For example, given three players $\{1, 2, 3\}$ and ratings r_1, r_2, r_3 such that player 1 has expected score $f(r_1 - r_2)$ against player 2, and player 2 has expected score $f(r_2 - r_3)$ against player 3, is it true that player 1 has expected score $f(r_1 - r_3)$ against player 3? The outcome of an experiment to test this hypothesis might depend on the particular players and what game they are playing. In practice our assumptions are computationally convenient and we expect that they will be a reasonable approximation to the truth.

There is some evidence [1] that the choice (4) of $f(z)$ (corresponding to the *Logistic* distribution) gives a better approximation to reality, at least for chess,

than the choice based on the normal distribution, as originally proposed by Elo [6]. Essentially, (4) is used in the current USCF rating system [7]. However, in this paper our choice of (4) is made primarily for computational convenience, and because it leads to an elegant algorithm.

By scaling the ratings, we can assume that $c = 1$. Thus, in the following we assume that

$$f(z) = \frac{1}{1 + e^{-z}} . \quad (5)$$

To avoid the exponential function, it is convenient to define

$$x_i = \exp(r_i) .$$

Thus, player i has expected score

$$\frac{1}{1 + x_j/x_i} = \frac{x_i}{x_i + x_j}$$

in a game against player j .

The ratings r_i are given by $r_i = \ln x_i$, but we can add an arbitrary constant β to all the r_i , since only their differences are significant (this corresponds to multiplying all the x_i by a positive constant $\kappa = \exp(\beta)$).

If $i \neq j$, the total number of games played between player i and player j is

$$g_{i,j} = s_{i,j} + s_{j,i} .$$

In the case $i = j$, we use this as a definition of $g_{i,i}$, that is $g_{i,i} = 2s_{i,i} = 2\sigma$.

Let $W = (w_{i,j})$ be a symmetric matrix of positive *weights* $w_{i,j}$. Games between players i and j will be weighted in proportion to $w_{i,j}$.

The *actual weighted score* of player i is

$$s_i = \sum_{j=1}^n s_{i,j} w_{i,j}$$

and, given the players' ratings, the *expected weighted score* is

$$e_i = \sum_{j=1}^n \left(\frac{x_i}{x_i + x_j} \right) w_{i,j} g_{i,j} .$$

If the only information available on the players' strengths is the results encoded in the matrix S , then it is reasonable to choose ratings such that the expected and actual weighted scores of each player are the same, that is

$$e_i = s_i \text{ for } i = 1, \dots, n .$$

Using the definitions of e_i , s_i and $g_{i,j}$, this condition is

$$\sum_{j=1}^n \left(\frac{x_i}{x_i + x_j} \right) w_{i,j} (s_{i,j} + s_{j,i}) = \sum_{j=1}^n s_{i,j} w_{i,j} \text{ for } i = 1, \dots, n . \quad (6)$$

2 Choosing Weights to Give a Linear Problem

The system of equations (6) is nonlinear in the unknowns x_i , and it is not immediately obvious that it has a solution, or how such a solution should be found. It is obvious that a solution is not unique, because if $x = (x_i)$ is one solution then so is κx for any positive constant κ .

In order to obtain a linear problem, we choose

$$w_{i,j} = (x_i + x_j)u_{i,j} \tag{7}$$

for some symmetric positive matrix U .

With the choice (7) of weights, equation (6) reduces to

$$\sum_{j=1}^n x_i u_{i,j} (s_{i,j} + s_{j,i}) = \sum_{j=1}^n s_{i,j} (x_i + x_j) u_{i,j} .$$

Using the symmetry of U , this simplifies to

$$x_i \sum_{j=1}^n a_{j,i} = \sum_{j=1}^n a_{i,j} x_j , \tag{8}$$

where $a_{i,j} = s_{i,j}u_{i,j}$. The matrix $A = (a_{i,j})$ is a weighted version of the score matrix S , and has the same sparsity pattern as S .

Let

$$d_i = \sum_{j=1}^n a_{j,i} .$$

d_i can be interpreted as the (weighted) number of points lost by player i , that is the (weighted) number of points scored by player i 's opponents in the games they played against i .

If a player does not lose any points, then the data is insufficient to determine a finite rating for him – he is “infinitely good”. Thus, we assume that $d_i > 0$ for $1 \leq i \leq n$.

Let $D = \text{diag}(d_i)$ be the diagonal matrix with diagonal elements d_i . By our assumption, D is nonsingular.

3 The Eigenvalue Problem

The condition (8) can be written in matrix-vector form as

$$Dx = Ax \tag{9}$$

or

$$D^{-1}Ax = x . \tag{10}$$

Thus, the solution vector x is the eigenvector corresponding to the eigenvalue 1 of the matrix $D^{-1}A$.

We observe that the matrix $A - D$ has linearly dependent rows; in fact, it is easy to see from the definition of D that the rows of $A - D$ sum to zero. Thus, $A - D$ is singular, so $D^{-1}A - I$ is singular, and $D^{-1}A$ does in fact have an eigenvalue 1. Similarly for AD^{-1} .

Equation (9) can be interpreted in terms of a Markov chain. Let $y = Dx$ and $M^T = AD^{-1}$. Then M is the transition matrix of a Markov chain ($m_{i,j} \geq 0$ and $\sum_j m_{i,j} = 1$). The vector $y/\|y\|_1$ gives the stationary probability distribution, because $y^T M = y^T$, or equivalently $AD^{-1}y = y$. It follows from standard theory of Markov matrices that $\rho(D^{-1}A) = \rho(AD^{-1}) \leq 1$.

In certain degenerate cases we can not expect finite ratings to be defined by the data. We already assumed that $d_i > 0$. This is necessary, but not sufficient. If the players can be split into two disjoint nonempty sets such that players in the first set always beat the players in the second set, then the players in the first set are “infinitely better” than the players in the second set. Similarly, if players in the first set never play players in the second set, we can not expect to compare their playing strengths. In practice, in either of these situations, we could split the problem and rate players in each set separately.

In the typical nondegenerate case, $D^{-1}A$ has a simple eigenvalue $\lambda_1 = 1$, and the other eigenvalues λ_i are inside the unit circle, that is $|\lambda_i| < 1$ except for $\lambda_1 = 1$. Then the *power method* converges and we can find x by the simple iteration

$$x^{(k+1)} = D^{-1}Ax^{(k)},$$

with a suitable starting vector, e.g. $x^{(0)} = (1, 1, \dots, 1)^T$.

So far we did not mention the role of the constant σ (recall that $s_{i,i} = \sigma$). The solution x of (10) is independent of σ , but the speed of convergence of the power method depends on σ . We have found in our experiments that $\sigma \in [0.2, 0.5]$ is a good choice to maximise the speed of convergence. Any $\sigma > 0$ will ensure that D is nonsingular.

4 Modifying the Weights

We have seen that solving an eigenvalue problem allows us to compute ratings if the score matrix is weighted by the weight function (7). It would be more natural to solve the problem with unit weights, that is $w_{i,j} = 1$. Unit weights have the advantage that, when applied to a round-robin (“all play all”) tournament, players with the same score obtain the same ratings, as is easy to see from (6).

The condition $w_{i,j} = 1$ is equivalent to

$$u_{i,j} = \frac{1}{x_i + x_j}.$$

Thus, we can regard the solution of the eigenvalue problem (10) as an inner iteration, and introduce an outer iteration where we change the weights. If $x^{(k)}$ is the solution to the k -th eigenvalue problem (with $x^{(0)} = (1, 1, \dots, 1)^T$ an

initial vector), then the $(k + 1)$ -st eigenvalue problem will use

$$u_{i,j}^{(k+1)} = \frac{1}{x_i^{(k)} + x_j^{(k)}}.$$

If the outer iteration converges, then it solves the original problem with weights

$$w_{i,j} = (x_i + x_j)u_{i,j} = 1.$$

In practice, we have found that convergence is quite rapid in nondegenerate cases. However, it is wasteful to solve the inner eigenvalue problems accurately. It is much more efficient to perform just one iteration of the power method in the inner loop. The resulting Algorithm 1 (without improvements to take advantage of sparsity) is given below.

```

for i := 1..n do
  x[i] := 1.0;
end for;
for k := 1, 2, ... until convergence do
  for i := 1..n do
    d[i] := 0.0;
  end for;
  for i := 1..n do
    sum := 0.0;
    for j := 1..n do
      temp := s[i,j]/(x[i] + x[j]);
      sum := sum + temp*x[j];
      d[j] := d[j] + temp;
    end for;
    y[i] := sum;
  end for;
  for i := 1..n do
    x[i] := y[i]/d[i];
  end for;
end for;

```

Algorithm 1: Unit Weights

Since the aim is to compute ratings $r_i = \ln x_i$, the convergence test should ensure a small relative error in each component of x . Thus, an appropriate stopping criterion is

$$\max_{1 \leq i \leq n} \left| \frac{x_i^{(k)} - x_i^{(k-1)}}{x_i^{(k)}} \right| < \varepsilon,$$

where ε is a tolerance depending on the accuracy required.

Failure to converge in a reasonable number of iterations may indicate that the problem is degenerate and that some ratings are diverging to $\pm\infty$. In this

case one or more players should in principal be excluded from consideration. A more convenient solution in practice is to add a “dummy” player who draws with all the other players, and whose games are given a positive weight γ , for example $\gamma = 1$. As $\gamma \rightarrow 0+$ the ratings tend to the correct values (possibly $\pm\infty$), but for any positive γ we obtain a nondegenerate problem and finite ratings¹.

If n is large then S (and hence A) will be sparse, since there are at most two off-diagonal entries for each game played. Thus, the number of nonzero elements is at most $2m + n$. The inner loop of the algorithm above essentially involves the multiplication of A on the right by x , and on the left by $[1, 1, \dots, 1]$. Thus, standard sparse matrix techniques can be used to reduce the complexity of the inner iteration from $O(n^2)$ to $O(m)$.

In practice the final ratings would be modified by a linear transformation to make them positive and not too small, before rounding to the nearest integer. (FIDE Elo ratings are usually in the range $[0, 3000]$, and BCF ratings are usually in the range $[0, 300]$, see [9].)

5 An Independent Derivation of Algorithm 1

From (6) with $w_{i,j} = 1$ we have, for $i = 1, \dots, n$,

$$\sum_{j=1}^n \left(\frac{x_i}{x_i + x_j} \right) (s_{i,j} + s_{j,i}) = \sum_{j=1}^n s_{i,j}. \quad (11)$$

This simplifies to

$$\sum_{j=1}^n \frac{x_i s_{j,i}}{x_i + x_j} = \sum_{j=1}^n \frac{x_j s_{i,j}}{x_i + x_j}$$

and, taking x_i outside the sum on the left, we see that

$$x_i = \left(\sum_{j=1}^n \frac{s_{i,j} x_j}{x_i + x_j} \right) / \left(\sum_{j=1}^n \frac{s_{j,i}}{x_i + x_j} \right). \quad (12)$$

A natural iteration to solve (12) is

$$x_i^{(k+1)} = \left(\sum_{j=1}^n \frac{s_{i,j} x_j^{(k)}}{x_i^{(k)} + x_j^{(k)}} \right) / \left(\sum_{j=1}^n \frac{s_{j,i}}{x_i^{(k)} + x_j^{(k)}} \right) \quad (13)$$

for $k = 1, 2, \dots$. However, it is easy to see that this is exactly the iteration implemented in Algorithm 1!

The significance of the diagonal terms $s_{i,i} = \sigma \geq 0$ is apparent if we consider the diagonal terms ($j = i$) in the numerator and denominator of (13). The

¹This solution is similar to the one adopted in the *PageRank* algorithm used by the Google search engine [2, 10, 11], where a fictional page essentially has links to every other page. Our parameter γ corresponds to the Page Rank algorithm’s $1 - d$.

diagonal term in the numerator is $\sigma/2$ and the diagonal term in the denominator is $\sigma/(2x_i^{(k)})$. As $\sigma \rightarrow \infty$ the right-hand side of (13) $\rightarrow x_i^{(k)}$. Thus, σ acts as a damping factor: larger values of σ tend to reduce the change $x_i^{(k)} \rightarrow x_i^{(k+1)}$ at each iteration.

Other iterations can be obtained in a similar manner. For example, taking x_i outside the sum on the left side of (11), we obtain the iteration

$$x_i^{(k+1)} = \left(\sum_{j=1}^n s_{i,j} \right) / \left(\sum_{j=1}^n \frac{s_{i,j} + s_{j,i}}{x_i^{(k)} + x_j^{(k)}} \right). \quad (14)$$

However, our numerical experiments suggest that (14) gives slower convergence than (13). This conclusion is confirmed by a first-order analysis in the special case that all the x_i are approximately equal.

6 Incorporating Old Ratings

Often some or all of the players will already have ratings, say player i has rating \hat{x}_i based on \hat{w}_i games². It is easy to take such “old” ratings into account by a slight modification of the argument leading to equation (6). We add $\hat{w}_i/2$ to the actual weighted score s_i , as if player i drew \hat{w}_i games against a player with rating \hat{x}_i , and also add $\hat{w}_i x_i / (x_i + \hat{x}_i)$ to the expected weighted score e_i . Thus equation (6) becomes

$$\sum_{j=1}^n \left(\frac{x_i}{x_i + x_j} \right) w_{i,j} (s_{i,j} + s_{j,i}) + \frac{\hat{w}_i x_i}{x_i + \hat{x}_i} = \sum_{j=1}^n s_{i,j} w_{i,j} + \frac{\hat{w}_i}{2} \quad (15)$$

for $i = 1, \dots, n$, and equation (11) becomes

$$\sum_{j=1}^n \left(\frac{x_i}{x_i + x_j} \right) (s_{i,j} + s_{j,i}) = \sum_{j=1}^n s_{i,j} - \left(\frac{x_i - \hat{x}_i}{x_i + \hat{x}_i} \right) \frac{\hat{w}_i}{2}. \quad (16)$$

Now, it is easy to see that the iteration (13) becomes

$$x_i^{(k+1)} = \frac{\left(\sum_{j=1}^n \frac{s_{i,j} x_j^{(k)}}{x_i^{(k)} + x_j^{(k)}} \right) + \frac{\hat{w}_i \hat{x}_i}{2(x_i^{(k)} + \hat{x}_i)}}{\left(\sum_{j=1}^n \frac{s_{j,i}}{x_i^{(k)} + x_j^{(k)}} \right) + \frac{\hat{w}_i}{2(x_i^{(k)} + \hat{x}_i)}} \quad (17)$$

for $k = 1, 2, \dots$

²The weight \hat{w}_i associated with an old rating might be reduced by a constant factor, say 0.5, to give less weight to old games than to recent ones.

7 Conclusion

We have shown how the computation of ratings for players of chess (and other games) can be reduced to an eigenvalue problem, or a sequence of eigenvalue problems, and that these eigenvalue problems are closely related to the problem of computing the stationary probability distributions of certain Markov chains. The eigenvalue problems can be solved efficiently by the power method. We derived an algorithm (Algorithm 1) by two different methods; one derivation (§§3–4) used the power method, the other (§5) was from first principles.

We have tested Algorithm 1 on small examples and on some larger examples with simulated scores. It would be interesting to test the algorithm on real data and to see how it compares with the methods currently in use for chess [9, 1, 6, 7] and other games, e.g. football [4]. Because they are practical systems that have evolved over time, most of these methods involve various *ad hoc* features and piecewise linear approximations, so they are less elegant (though perhaps more practical) than the relatively simple algorithm discussed here.

Acknowledgement

Thanks to my friend and mentor Gene Golub who persuaded me to write up this work, and commented on an early draft. Unfortunately, Gene passed away in 2007.

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